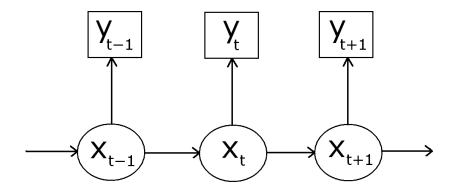
Time Series and Dynamic Models

Section 4 The AR(1) + noise model Our first DLM...

Carlos M. Carvalho The University of Texas at Austin

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In general:



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The Model:

$$y_t = \theta_t + \nu_t$$

$$\theta_t = \alpha + \beta \theta_{t-1} + \epsilon_t$$

•
$$u_t \sim N(0, \tau^2)$$
 and $\epsilon_t \sim N(0, \omega^2)$... independent.

- $p(\theta_0|D_0) = N(m_0, C_0)$ This is the posterior for θ_0 at time 0
- $D_t = \{D_{t-1}, y_t\}$ is the information set up to t.
- Assume, for now, knowledge of α , β , ω^2 and τ^2
- Is this model weakly stationary?

1. Prior for θ_1 :

$$p(\theta_1|D_0) = \int p(\theta_1|\theta_0, D_0) p(\theta_0|D_0) d\theta_0$$

= $N(a_1, R_1)$

where

$$a_1 = lpha + eta m_0$$
 and $R_1 = eta^2 C_0 + \omega^2$

2. Predictive for y_1 :

$$p(y_1|D_0) = \int p(y_1|\theta_1, D_0)p(\theta_1|D_0)d\theta_1$$
$$= N(f_1, Q_1)$$

where

$$f_1 = a_1$$
 and $Q_1 = R_1 + au^2$

3. Posterior for θ_1 :

$$p(\theta_1|D_1) \propto p(y_1|\theta_1, D_0)p(\theta_1|D_0)$$

$$\begin{bmatrix} \theta_1 \\ y_1 \end{bmatrix} \sim N\left(\begin{bmatrix} a_1 \\ f_1 \end{bmatrix}, \begin{pmatrix} R_1 & R_1 \\ R_1 & Q_1 \end{pmatrix} \right)$$
(Why?) Therefore,

$$[\theta_1|D_1] \sim N(m_1, C_1)$$

where

$$m_1 = a_1 + A_1 e_1$$

 $C_1 = R_1 - R_1 A_1$

with

$$A_1=R_1/Q_1$$
 and $e_1=(y_1-f_1)$

• What about $(\theta_0|D_1)$? (what does that mean)?

$$p(\theta_0|D_1) = \int p(\theta_0|\theta_1, D_1)p(\theta_1|D_1)d\theta_1$$
$$= \int p(\theta_0|\theta_1, D_0)p(\theta_1|D_1)d\theta_1$$

where

$$p(heta_0| heta_1,D_0) \propto p(heta_1| heta_0,D_0)p(heta_0|D_0)$$

...okay, in this model we should be able to solve the above integral... an easier way is to work with the multivariate normal representation:

$$\begin{bmatrix} \theta_0 \\ y_1 \end{bmatrix} \sim N\left(\begin{bmatrix} m_0 \\ f_1 \end{bmatrix}, \begin{pmatrix} C_0 & \beta C_0 \\ \beta C_0 & Q_1 \end{pmatrix}\right)$$

(Why?)

Therefore,

 $[\theta_0|D_1] \sim \textit{N}(\textit{h}_0,\textit{H}_0)$

where

$$h_0 = m_0 + \beta C_0 Q_1^{-1} e_1$$

$$H_0 = C_0 - \beta^2 C_0^2 Q_1^{-1}$$

Filter-Forward Recursions

We can generalize the above discussion by the following:

- Posterior at t 1: $[\theta_{t-1}|D_{t-1}] \sim N(m_{t-1}, C_{t-1})$
- Prior at t 1: $[\theta_t | D_{t-1}] \sim N(a_t, R_t)$ with

$$a_t = lpha + eta m_{t-1}$$
 and $R_t = eta^2 C_t + \omega^2$

▶ Predictive at t - 1: $[y_t | D_{t-1}] \sim N(f_t, Q_t)$ with

$$f_t = a_t$$
 and $Q_t = R_t + \tau^2$

• Posterior at t: $[\theta_t|D_t] \sim N(m_t, C_t)$

$$m_t = a_t + A_t e_t$$
 and $C_t = R_t - R_t A_t$

with

$$A_t = R_t/Q_t$$
 and $e_t = y_t - f_t$



- What about α , β , ω^2 and τ^2 ?
- Could we handle

$$[\alpha, \beta, \omega^2, \tau^2 | \theta_{1:T}, D_T]$$

If so, a Gibbs sampler iterates through (drawing)
 1. [α, β, ω², τ²|θ_{1:T}, D_T]
 2. [θ_{1:T}|α, β, ω², τ², D_T]

Our First FFBS

FFBS stands for *Filter Forward Backward Sampling* This is what the Gibbs Sample for models like the AR(1) plus noise is called due to its form... we'll see below.

Our Goal: Obtain samples from the joint posterior of $(\alpha, \beta, \tau^2, \omega^2, \theta_{1:T})$

How: Build a Gibbs sampler that iterates through (drawing)

- 1. $p(\alpha, \beta, \omega^2, \tau^2 | \theta_{1:T}, D_T)$
- 2. $p(\theta_{1:T}|\alpha,\beta,\omega^2,\tau^2,D_T)$
- Step 1 should be easy, right?
- Step 2 requires some work...

$$p(\theta_{1:T}|\alpha,\beta,\omega^2,\tau^2,D_T)$$

For notation simplicity, let us drop the fixed parameters from the conditioning set, i.e, $p(\theta_{1:T}|\alpha, \beta, \omega^2, \tau^2, D_T) = p(\theta_{1:T}|D_T)$

$$p(\theta_{1:T}|D_T) = p(\theta_1, \theta_2, \dots, \theta_T | \alpha, \beta, \omega^2, \tau^2, D_T)$$

= $p(\theta_1|\theta_2, \dots, \theta_T, D_T)p(\theta_2|\theta_3, \dots, \theta_T, D_T) \dots p(\theta_T|D_T)$
= $\prod_{t=1}^{T-1} p(\theta_t|\theta_{t+1:T}, D_T)p(\theta_T|D_T)$

- ► So, we know $p(\theta_T | D_T)$, right? $p(\theta_T | D_T) = N(m_T, C_T)$
- How about $p(\theta_t | \theta_{t+1:T}, D_T)$?

$$p(\theta_{1:T}|\alpha,\beta,\omega^2,\tau^2,D_T)$$

Given the conditional independence structure of the model we can write

$$p(\theta_t | \theta_{t+1:T}, D_T) = p(\theta_t | \theta_{t+1}, D_t)$$

(Why?)

Okay, now this should be straightforward as

 $p(\theta_t|\theta_{t+1}, D_t) \propto p(\theta_{t+1}|\theta_t, D_t)p(\theta_t|D_t)$

$$\begin{bmatrix} \theta_t \\ \theta_{t+1} \end{bmatrix} \sim N\left(\begin{bmatrix} m_t \\ a_{t+1} \end{bmatrix}, \begin{pmatrix} C_t & \beta C_t \\ \beta C_t & R_{t+1} \end{pmatrix}\right)$$

 $p(\theta_{1:T}|\alpha,\beta,\omega^2,\tau^2,D_T)$

Hence,

$$p(\theta_t|\theta_{t+1}, D_t) = N(h_t, B_t)$$

where

$$h_t = m_t + \beta C_t / R_{t+1} (\theta_{t+1} - a_{t+1}) B_t = C_t - \beta^2 C_t^2 / R_{t+1}$$

Now what?

- Okay, now what??
- It looks like we can, conditional on the fixed parameters defining the system, filter forward and get to p(θ_T|D_T)
- We then draw $\theta_T^{(1)}$ from $p(\theta_T | D_T)$...
- Now, we should be able to sample $\theta_{T-1}^{(1)}$ from $p(\theta_{T-1}|\theta_T^{(1)}D_{T-1})...$ keep going until we get to...
- $\{\theta_1^{(1)}, \theta_2^{(1)}, \dots, \theta_{T-1}^{(1)}, \theta_T^{(1)}\}\$ a draw from the joint distribution $p(\theta_{1:T}|\alpha, \beta, \omega^2, \tau^2, D_T)$

Homework

- Estimate an AR(1) plus noise model to the "daily temperature in Austin data" available in the class website. Only use the last 1,000 observations.
- Report histograms of your posterior distribution for the fixed parameters
- Plot the time series of the observed temperature along with the posterior mean and posterior 95% range for each latent state.
- Predict the temperature for the next 20 days... plot the predictions and prediction intervals.